

**Definition 1** (Statistical Experiment). A statistical experiment is a triple  $(\mathcal{X}, \mathcal{F}, \{\mathbf{P}_\theta\}_{\theta \in \Theta})$  where  $\mathcal{X}$  is the sample space,  $\mathcal{F}$  is a  $\sigma$ -algebra on  $\mathcal{X}$ , and  $\{\mathbf{P}_\theta\}_{\theta \in \Theta}$  is a collection of probability measures on  $\mathcal{X}$  parametrized by  $\theta$  in the parameter space  $\Theta$ .

**Definition 2** (Sufficient Statistic). Let  $(\mathcal{X}, \mathcal{F}, \{\mathbf{P}_\theta\}_{\theta \in \Theta})$  be an statistical experiment. Let a measurable function  $T : (\mathcal{X}, \mathcal{F}) \rightarrow (\mathcal{T}, \mathcal{G})$  be a statistic. The statistic  $T$  is said to be sufficient if  $\mathbf{P}_\theta(\cdot | T)$  does not depend on  $\theta$ .

**Example 1.** Consider an experiment  $(\mathcal{X}, \mathcal{F}, \{\mathbf{P}_\theta\}_{\theta \in [0,1]})$  where  $\mathcal{X} = \{0, 1\}^n$ ,  $\mathcal{F}$  is the power set of  $\mathcal{X}$ , and  $\mathbf{P}_\theta$  is the joint distribution of  $n$  iid Bernoulli( $\theta$ ) distributions. Define statistic  $T(x) := \sum_{i=1}^n x_i$  where  $x_i$  denotes the  $i$ -th component of  $x$ . We check that  $T$  is a sufficient statistic: Given any  $x \in \mathcal{X}$ , let  $t := T(x)$  and we have

$$\mathbf{P}_\theta\{X = x | T = t\} = \frac{\mathbf{P}_\theta\{X = x\}}{\mathbf{P}_\theta\{T = t\}} = \frac{\prod_{i=1}^n \theta^{x_i} (1 - \theta)^{1-x_i}}{\binom{n}{t} \theta^t (1 - \theta)^{n-t}} = \binom{n}{t}^{-1}.$$

Since the conditional distribution does not depend on  $\theta$ ,  $T$  is a sufficient statistic.

**Theorem 1** (Fisher-Neyman Factorization). Consider a statistical experiment  $(\mathcal{X}, \mathcal{F}, \{\mathbf{P}_\theta\}_{\theta \in \Theta})$ . A statistic  $T : (\mathcal{X}, \mathcal{F}) \rightarrow (\mathcal{T}, \mathcal{G})$  is a sufficient statistic iff there exists measurable functions  $\{g_\theta\}_{\theta \in \Theta}$  defined on  $(\mathcal{T}, \mathcal{G})$  and  $h$  defined on  $(\mathcal{X}, \mathcal{F})$  such that the probability density function can be decomposed as

$$f(x | \theta) = g_\theta(T(x))h(x). \quad (1)$$

*Proof.* We only show the proof for discrete random variables. <sup>1</sup> In this case, we have the probability density function  $f(x | \theta) = \mathbf{P}_\theta\{X = x\}$ .

( $\Rightarrow$ ) Suppose that  $T$  is sufficient. Given any realization  $x \in \mathcal{X}$ , let  $t := T(x)$  and we have the following

$$\begin{aligned} \mathbf{P}_\theta\{X = x\} &= \mathbf{P}_\theta\{X = x, T = t\} \\ &= \mathbf{P}_\theta\{X = x | T = t\} \mathbf{P}_\theta\{T = t\} \\ &= \underbrace{\mathbf{P}\{X = x | T = t\}}_{h(x)} \underbrace{\mathbf{P}_\theta\{T = t\}}_{g_\theta(t)} \end{aligned}$$

where the last equality is obtained by the definition of sufficiency.

( $\Leftarrow$ ) Suppose we have the factorization (1). Given any realization  $x \in \mathcal{X}$ , let  $t := T(x)$  and we have

$$\begin{aligned} \mathbf{P}_\theta\{X = x | T = t\} &= \frac{\mathbf{P}_\theta\{X = x, T = t\}}{\mathbf{P}_\theta\{T = t\}} \\ &= \frac{g_\theta(t)h(x)}{\mathbf{P}_\theta\{T = t\}} \end{aligned}$$

<sup>1</sup>For continuous random variables, one must deal with measure-theoretic technicalities in the proof. One can find a proof for general random variables in *Chapter 2.2 Statistics and Sufficiency* from [1].

where the second equality is obtained by using the factorization on the numerator. Now consider the denominator, we have

$$\mathbf{P}_\theta\{T = t\} = \sum_{x':T(x')=t} \mathbf{P}_\theta\{X = x'\} = \sum_{x':T(x')=t} g_\theta(t)h(x')$$

where the second equality is obtained by factorization. Substitute the result back and we have

$$\begin{aligned} \mathbf{P}_\theta\{X = x | T = t\} &= \frac{g_\theta(t)h(x)}{\sum_{x':T(x')=t} g_\theta(t)h(x')} \\ &= \frac{h(x)}{\sum_{x':T(x')=t} h(x')} \end{aligned}$$

where the final expression does not depend on  $\theta$ . Hence  $T$  is a sufficient statistic for  $\theta$ . #

**Example 2** (Example 1 Continued). Now we check that  $T(x) = \sum_{i=1}^n x_i$  is a sufficient statistic using [Theorem 1](#):

$$\mathbf{P}_\theta\{X = x\} = \prod_{i=1}^n \theta^{x_i} (1 - \theta)^{1-x_i} = \underbrace{\theta^{T(x)} (1 - \theta)^{n-T(x)}}_{g_\theta(T(x))}.$$

In this case,  $h(x) = 1$  is a constant function. Hence,  $T$  is a sufficient statistic.

**Definition 3** (Sufficiency Principle). Let  $(\mathcal{X}, \mathcal{F}, \{\mathbf{P}_\theta\}_{\theta \in \Theta})$  be a statistical experiment and let  $T$  be a sufficient statistic. The sufficiency principle states that inference about  $\theta$  should depend only on  $T$ . That is, if two samples  $x$  and  $x'$  satisfies  $T(x) = T(x')$ , then they should lead to the same inference about  $\theta$ .

**Remark 1.** The concept of sufficient statistic and the sufficiency principle are both due to Sir Ronald Fisher in the early 20th century.

## References

- [1] Jun Shao. *Mathematical Statistics*. New York: Springer, 1999. ISBN: 0-387-98674-X.